

On the Bayesian formulation of fractional inverse problems and data-driven discretization of forward maps

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- 1 Bayesian formulation of fractional inverse problems.
- 2 Data driven discretization of forward maps.

This presentation mostly based on:

- ① *The Bayesian Formulation and Well-Posedness of Fractional Elliptic Inverse Problems* (2017 Inverse Problems) with D. Sanz-Alonso.
- ② *Data driven discretizations of forward maps in Bayesian inverse problems* (In preparation) with D. Bigoni, Y. Marzouk and D. Sanz-Alonso.

Part 1: Bayesian formulation of fractional inverse problems.

Inverse problem: learn a permeability field from partial and noisy observations of pressure field.

PDE version: Learn diffusion coefficient and **order** of a (FPDE) based on partial and noisy observations of its solution.

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$$u = (s, A) \longrightarrow \mathcal{F}(u) \longrightarrow \mathcal{O} \circ \mathcal{F}(u)$$
$$\mathcal{G} := \mathcal{O} \circ \mathcal{F}$$

$$y = \mathcal{G}(u) + \text{Noise}$$

Inverse problem: learn a permeability field from partial and noisy observations of pressure field.

PDE version: Learn diffusion coefficient and **order** of a (FPDE) based on partial and noisy observations of its solution.

$$u = (s, A) \longrightarrow \mathcal{F}(u) \longrightarrow \mathcal{O} \circ \mathcal{F}(u)$$
$$\mathcal{G} := \mathcal{O} \circ \mathcal{F}$$

$$\phi(y; \mathcal{G}(u))$$

- **Forward map:** $p = \mathcal{F}(u)$.

$$\begin{cases} L_{AP}^s p = f, & \text{in } D, \\ \partial_{AP} p = 0, & \text{on } \partial D, \end{cases} \quad (1)$$

where $\partial_{AP} p := A(x)\nabla p \cdot \nu$, and ν is the exterior unit normal to ∂D .

- **Observation map:** $\mathcal{O}(p) := (p(x_1), \dots, p(x_n))$ for some $x_i \in D$.
- **Noise model:** $\phi(y, \mathcal{G}(u)) = \exp\left(-\frac{1}{2\gamma^2}\|y - \mathcal{G}(u)\|^2\right)$.

- **Forward map:** $p = \mathcal{F}(u)$.

$$\begin{cases} L_A^s p = f, & \text{in } D, \\ \partial_A p = 0, & \text{on } \partial D, \end{cases} \quad (2)$$

where $\partial_A p := A(x)\nabla p \cdot \nu$, and ν is the exterior unit normal to ∂D .

Here,

$$L_A^s p = \sum_{k=1}^{\infty} \lambda_{A,k}^s p_k \psi_{A,k}.$$

- **Observation map:** $\mathcal{O}(p) := (p(z_1), \dots, p(z_m))$ for some $z_i \in D$.
- **Noise model:** $\phi(y, \mathcal{G}(u)) = \exp\left(-\frac{1}{2\gamma^2} \|y - \mathcal{G}(u)\|^2\right)$.

- A. M. Stuart. Inverse problems: a Bayesian perspective. (2010).
- J. Kaipio and E. Somersalo. Statistical and computational inverse problems (2006).

- Prior: $u \sim \pi_u$
- Likelihood model: $\pi_{y|u}$
- Bayes rule (informally):

$$\nu^y(u) := \pi_{u|y} \propto \pi_{y|u} \cdot \pi_u$$

Posterior distribution.

ν^y is the fundamental object in Bayesian inference.

- Estimates:

$$\mathbb{E}_{u \sim \nu^y} (R(u))$$

- Uncertainty quantification:

$$\text{Var}_{u \sim \nu^y} (R(u))$$

Advantages of Bayesian formulation

Well defined mathematical framework:

- Stability (Well-posedness).
- Posterior consistency (contraction rates, scalings for parameters, etc).
- Consistency of numerical methods.

- π_u is a distribution on $(0, 1) \times \mathcal{H}$.
- For example,

$$\pi_u = \pi_s \otimes \pi_A$$

$$A = e^v I_d, \quad \text{where } v \sim N(0, K)$$

- Karhunen-Loeve expansion:

$$v = \sum_{i=1}^{\infty} \lambda_{K,i} \zeta_i \psi_i, \quad \zeta_i \sim N(0, 1).$$

Theorem (NGT and D. Sanz-Alonso 17')

Suppose that \mathcal{G} is continuous in $\text{supp}(\pi_u)$. Then, posterior distribution ν^y is absolutely continuous with respect to prior:

$$d\nu^y(u) \propto \phi(y; \mathcal{G}(u))d\pi_u(u),$$

Recall: $\mathcal{G} : (s, A) \rightarrow \mathbb{R}^m$.

Theorem (NGT and D. Sanz-Alonso 17')

Suppose that $\mathcal{G} \in L^2_{\pi_u}$. Then the map

$$y \mapsto \nu^y$$

is Locally Lipschitz in the Hellinger distance. That is, For $|y_1|, |y_2| \leq r$ we have

$$d_{hell}(\nu^{y_1}, \nu^{y_2}) \leq C_r \|y_1 - y_2\|.$$

The analysis reduces to studying stability through regularity of FPDEs.

- L. A. Caffarelli and P. R. Stinga. Fractional elliptic equations, Caccioppoli estimates and regularity. (2016)

- M. Dashti and A. M. Stuart. Uncertainty quantification and weak approximation of an elliptic inverse problem. (2011).
- S. Agapiou, S. Larsson, and A.M. Stuart. Posterior contraction rates for the Bayesian approach to linear ill-posed inverse problems. (2013) .
- S. Volmer. Posterior consistency for Bayesian inverse problems through stability and regression results. (2013).

- Need a way to approximate expectations with respect to ν^y .
- Standard procedure: MCMC.
Generate a path of a Markov chain with invariant distribution $\nu^y: u_1, \dots, u_k, \dots$ and then use

$$\frac{1}{k} \sum_{i=1}^k R(u_i)$$

- However, careful with:
 - 1 Discretization of u .
 - 2 Discretization of forward map.

For the sake of simplicity assume $A = e^{\nu} \cdot I_d$ and known $s \in (0, 1)$.

Metropolis Hastings with pCN proposal:

Having defined v_k , v_{k+1} is generated according to:

- 1 Proposal: $\tilde{v} = \sqrt{1 - \beta^2} v_k + \beta \xi$, where $\xi \sim \pi_{\nu}$.
- 2 Acceptance probability:

$$\alpha(\tilde{u}, u_k) := \min \left\{ 1, \frac{\phi(y; \mathcal{G}(\tilde{u}))}{\phi(y; \mathcal{G}(u_k))} \right\}$$

- 3 $v_{k+1} := \begin{cases} \tilde{v} & \text{with prob } \alpha(\tilde{u}, u_k) \\ v_k & \text{with prob } 1 - \alpha(\tilde{u}, u_k) \end{cases}$

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Compare to: $\tilde{v} = v_k + \beta \xi$

S. L. Cotter, G. O. Roberts, A. M. Stuart, and D. White. MCMC methods for functions: modifying old algorithms to make them faster. Statistical Science.

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Idealized MCMC algorithm

For the sake of simplicity assume $A = e^\nu \cdot I_d$ and known $s \in (0, 1)$.

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S. L. Cotter, G. O. Roberts, A. M. Stuart, and D. White. MCMC methods for functions: modifying old algorithms to make them faster. Statistical Science.

Robustness to truncation:

$$\xi = \sum_{i=1}^L \lambda_{K,i} \zeta_i \Psi_i, \quad \zeta_i \sim N(0, 1)$$

- 2 Acceptance probability:

$$\alpha(\tilde{u}, u_k) := \min \left\{ 1, \frac{\phi(y; \mathcal{G}(\tilde{u}))}{\phi(y; \mathcal{G}(u_k))} \right\}$$

Part 2: Data driven discretization of forward maps

$$u \longrightarrow \mathcal{F}(u) \longrightarrow \mathcal{O} \circ \mathcal{F}(u)$$
$$\mathcal{G} := \mathcal{O} \circ \mathcal{F}$$

$$\phi(y; \mathcal{G}(u))$$

To produce u_{k+1} :

- 1 Proposal: $\tilde{u} = \sqrt{1 - \beta^2} u_k + \beta \xi$, $\xi \sim \pi_u$.
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How do we choose the discretization? How fine?
Inhomogeneous in space?

$$(u, X) \longrightarrow \mathcal{F}_X(u) \longrightarrow \mathcal{O} \circ \mathcal{F}(u)$$
$$\mathcal{G}_X := \mathcal{O} \circ \mathcal{F}_X$$

$$\phi(y; \mathcal{G}_X(u))$$

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More specifically: $X = (N, \{x_1, \dots, x_N\})$.

True and Surrogate Bayesian inverse problems

Prior: $u \sim \pi_u$.

Likelihood: $\phi(y; \mathcal{G}(u))$

Posterior:

$$d\nu^y(u) \propto \phi(y; \mathcal{G}(u))d\pi_u(u)$$

Prior: $(u, X) \sim \pi_{u,X}$.

Likelihood: $\phi(y; \mathcal{G}_X(u))$

Posterior:

$$d\nu^y(u, X) \propto \phi(y; \mathcal{G}_X(u))d\pi_{u,X}(u, X)$$

Prior for surrogate problem

For simplicity our prior takes the form:

$$\pi_{u,X} = \pi_{x_1, \dots, x_N|N} \cdot \pi_N \cdot \pi_u.$$

- π_u is as for the true problem.
- $\pi_{u,X}$ treats X and u independently.
- $\pi_{x_1, \dots, x_N|N} = dx_1 \dots dx_N$ on D^N .
- π_N takes into account cost of discretization of \mathcal{F} using N elements:

$$\pi_N \propto \exp(-C(N))$$

$$\phi(y; \mathcal{G}_X(u))$$

Recall $\mathcal{G}_X = \mathcal{O} \circ \mathcal{F}_X$.

$$\phi(y; \mathcal{G}_X(u))$$

Recall $\mathcal{G}_X = \mathcal{O} \circ \mathcal{F}_X$.

However, we don't triangulate directly using X . First, we regularize the points x_1, \dots, x_n .

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Recall $\mathcal{G}_X = \mathcal{O} \circ \mathcal{F}_X$.

However, we don't triangulate directly using X . We regularize the points x_1, \dots, x_n .

$X = (\{x_1, \dots, x_n\}, N)$ induces a density estimator ρ_X .

- 1 ρ_X is used to choose points in a master grid.

or

- 2 Using ρ_X we start a flow of the points x_1, \dots, x_n attempting to minimize an energy of the form:

$$E(x_1, \dots, x_n) \sim \sum_{i,j} \exp(-|x_i - x_j|^2 / (h_N \cdot \rho_X(x_i))^2)$$

Metropolis within Gibbs:

Alternate:

- 1 Update the unknown; discretization is fixed. $u|X, y$
- 2 Change distribution of elements. $x_1, \dots, x_N|N, u, y$
- 3 Coarsen or refine discretization. $x_1, \dots, x_N, N|u, y$

Given (u_k, X_k) produce u_{k+1} by:

- 1 Proposal: $\tilde{u} = \sqrt{1 - \beta^2} u_k + \beta \xi$, $\xi \sim \pi_u$.
- 2 Compute acceptance probability:

$$\alpha(\tilde{u}, u_k) := \min \left\{ 1, \frac{\phi(y; \mathcal{G}_{X_k}(\tilde{u}))}{\phi(y; \mathcal{G}_{X_k}(u_k))} \right\}$$

- 3 $u_{k+1} := \begin{cases} \tilde{u} & \text{with prob } \alpha(\tilde{u}, u_k) \\ u_k & \text{with prob } 1 - \alpha(\tilde{u}, u_k) \end{cases}$

Change distribution of elements

Given (u_k, X_k) produce $x_1^{k+1}, \dots, x_N^{k+1}$ by:

- 1 Proposal: $\tilde{x}_i = x_i + F(x_i)$, for all $i = 1, \dots, N$.

$$F := (\Psi_1, \Psi_2) \sim N(0, K_1 \otimes K_2)$$

- 2 Compute acceptance probability:

$$\alpha(\tilde{X}, X_k) := \min \left\{ 1, \frac{\phi(y; \mathcal{G}_{\tilde{X}}(u_k))}{\phi(y; \mathcal{G}_{X_k}(u_k))} \cdot \frac{p(X_k | \tilde{X})}{p(\tilde{X} | X_k)} \right\}$$

- 3 $X_{k+1} := \begin{cases} \tilde{X} & \text{with prob } \alpha(\tilde{X}, X_k) \\ X_k & \text{with prob } 1 - \alpha(\tilde{X}, X_k) \end{cases}$

Coarsen or refine discretization

Given (u_k, X_k) we produce N_{k+1} and $x_1^{k+1}, \dots, x_{N_{k+1}}^{k+1}$ by:

- 1 Proposal: First construct ρ_X and generate $\tilde{N} \sim p(\cdot | N_k)$.
 - If $\tilde{N} < N$ let $\tilde{x}_i = x_i$ for $i = 1, \dots, \tilde{N}$.
 - If $\tilde{N} \geq N$ let $\tilde{x}_i = x_i$ for $i = 1, \dots, N$ and generate $\tilde{x}_{N+j} \sim \rho_X$ for $j = 1, \dots, \tilde{N} - N$.
- 2 Compute acceptance probability:

$$\alpha(\tilde{X}, X_k) := \min \left\{ 1, \frac{\phi(y; \mathcal{G}_{\tilde{X}}(u_k))}{\phi(y; \mathcal{G}_{X_k}(u_k))} \cdot p(X, \tilde{X}) \cdot \frac{p(\tilde{N} | N_k)}{p(N_k | \tilde{N})} \cdot \frac{\exp(-C(\tilde{N}))}{\exp(-C(N))} \right\} \quad (3)$$

- 3 $X_{k+1} := \begin{cases} \tilde{X} & \text{with prob } \alpha(\tilde{X}, X_k) \\ X_k & \text{with prob } 1 - \alpha(\tilde{X}, X_k) \end{cases}$

Conclusions and future work

- 1 Bayesian formulation of fractional inverse problems.
- 2 Data driven discretization of forward maps.

Thank you for your attention!